

Smart or Smash? The Effect of Financial Sanctions on Trade in Goods and Services

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Impressum:

CESifo Working Papers

ISSN 2364-1428 (electronic version)

Publisher and distributor: Munich Society for the Promotion of Economic Research - CESifo GmbH

The international platform of Ludwigs-Maximilians University's Center for Economic Studies and the ifo Institute

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Editor: Clemens Fuest

<https://www.cesifo.org/en/wp>

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Abstract

We examine the extent to which financial sanctions imposed by Germany through its European Union and United Nations commitments cause collateral damage on Germany's trade in goods and services. Financial sanctions reduce Germany's inflows and outflows of financial assets, as well as imports and exports of goods and services. The relative effects on trade in goods and services are weaker than on financial assets, about half as large in the case of goods and two-thirds as large in the case of services. The effect on trade in goods is entirely due to episodes where financial sanctions are accompanied by export restrictions of specific goods. In the case of services trade, only exports are affected by financial sanctions once export restrictions are considered. The primary channel through which sanctions affect the three types of cross-border flows is the extensive margin. Anticipation effects are quite strong for financial assets and weak for services and goods.

JEL-Codes: F200, F360, F380.

Keywords: sanction, restriction, cross-border transaction, trade in goods, trade in services, financial flows.

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August 24, 2023

Forthcoming: Review of International Economics

We thank two anonymous referees, Ulrich Grosch, Stephan Jank, Axel Jochem, Jean-Francois Maystadt, Yoto Yotov and conference and seminar participants at the University of Bari Aldo Moro, Lund University, Purdue University, and at Deutsche Bundesbank for helpful comments. Discussion Papers represent the authors' personal opinions and do not necessarily reflect the views of the Deutsche Bundesbank or the Eurosystem.

1. Introduction

Sanctions have long been used as a foreign policy tool to achieve a variety of objectives. Dating back at least to the ancient Greeks and the Megarian decree of Athenians in 435 BC, sanctions have been in place throughout history, with various degrees of frequency and intensity, from U.S. sanctions of France and Great Britain during the Napoleonic wars to sanctions imposed on South Africa's apartheid regime in the 1980s and, most recently, sanctions against Russia and Iran (Drezner, 1999). While sanctions often vary in their mechanics and specific targets, their goal is usually similar – to inflict economic pain to force a change in policy. As documented by Felbermayr et al. (2020b) sanctions are either complete, trying to fully cutoff the target, or partial, aiming at a subset of activities or specific actors and their access to the global economy. The more recent cases of sanctions tend to be of the partial kind and policy makers have started referring to them as smart: targeting only specific activities and specific individuals, firms and organizations, thereby minimizing their spillover effects or collateral damage.

Apart from classifying sanctions as partial or complete, Felbermayr et al. (2020b) distinguish between six types of sanctions: trade, financial, travel restrictions, arms, military assistance, and other (which primarily entail diplomatic measures). The primary objective of each type of sanctions is obvious. Trade sanctions are designed to reduce trade between the sender country, the country imposing sanctions, and the target country, the country being sanctioned. As shown by Felbermayr et al. (2020b) in a large cross-country sample, trade sanctions do indeed reduce trade between the sender and the target. Similarly, financial sanctions reduce cross-border financial flows as shown by Besedeš et al. (2017). Travel restrictions limit travel between the two countries and are usually imposed by prohibiting specific individuals from entering the sender country. In addition to the direct consequences of sanctions, however, there is the possibility and open question of secondary effects, that the effect of one type of sanctions may spill over into another sphere of cross-border interactions and, therefore, cause collateral damage. Perhaps the most obvious connection is the possibility of financial sanctions to reduce trade flows between the target and sender country. After all, to conduct trade, money or finance must flow between countries.² If the flow of money is curtailed, it only stands to reason that there will be an accompanying reduction in trade. Indeed, Felbermayr et al. (2020b) find some evidence, albeit weak, of these spillover effects in their

² Apart from rare exception cases such as barter trade.

cross-country study where they show that trade between sender and target countries is reduced not only by trade sanctions, but also by financial sanctions.

Jäkel et al. (2022) identify four channels or mechanisms through which sanctions of any kind can affect cross-border flows. The channels are (i) the legal enforcement of sanctions, (ii) self-imposed restriction in light of sanctions; (iii) incurring information costs as a result of sanctions, and (iv) uncertainty about the evolution of sanctions. Financial sanctions can affect both trade and cross-border financial flows through all four channels. The legal enforcement channel has the most direct effect across all flows as it may affect the ability to finance cross-border flows. The remaining three channels can be thought of as providing second-order effects as they do not directly infringe on the ability to conduct cross-border flows, but do increase costs of doing business across borders or result in reduced willingness to engage in such business.

This paper extends our efforts to understand the effects of financial sanctions imposed by Germany based on it being both a member of the European Union (EU) and the United Nations (UN).³ In the first analysis we showed that financial sanctions reduce German cross-border financial flows by about 50% (Besedeš et al. 2017). In another exercise (Besedeš et al. 2021) we examined how precise, or how smart, these sanctions are in terms of which types of firms they affect. We focused on the performance of non-financial firms and showed that while financial sanctions reduce financial flows, they have no adverse effect on broader measures of firm performance such as employment or total sales. This may be because German non-financial firms affected by these sanctions tend to be disproportionately large and are able to expand their activities with non-sanctioned countries. In addition, Efung et al. (2023) reveal that banks located in Germany decrease external positions in sanctioned countries while branches and subsidiaries abroad do not respond. For affiliated banks located in countries with low financial standards, they even observe a relative increase in credit supply. Finally, Drott et al. (2022) examine the effect of financial sanctions against Russia at the most disaggregated level possible, individual bank accounts. Using data at daily frequency from the Eurosystem's real-time gross settlement system TARGET2, they provide empirical evidence that sanctions imposed by the EU on Russian banks following the country's military interventions in Ukraine in 2014 and 2022 have sizably reduced financial transactions with sanctioned Russian bank

³ Since European Union member countries are committed to the Common Foreign and Security Policy, any sanctions imposed by Germany have been imposed by the EU in the first place.

accounts. Among the various sanction measures taken, exclusion from SWIFT, a global provider of secure financial messaging services, turns out to have the largest effects.

In this paper we turn our attention to the question of whether financial sanctions affect international trade. We investigate the effect of financial sanctions on trade in goods and services, respectively. To properly gauge this secondary effect of financial sanctions, the spillover into other activities or collateral damage, we also re-examine the effect on German cross-border financial flows, allowing us to compare the magnitudes of the primary or direct effect and the secondary or the spillover effect. We do so by using four different data sets. One reflects financial sanctions imposed by Germany between 2001 and 2020; it is an update of data used by Besedeš et al. (2017) which reflected only the period between 2005 and 2014. This information is combined with three different data sets: data on merchandise trade sourced from Eurostat and data on trade in services as well as cross-border financial flows both sourced from the Deutsche Bundesbank. Our data span the period between January 2001 and September 2020. We also take advantage of one feature of financial sanctions imposed by Germany. While all 29 episodes of sanctions in our data are sanctions that freeze financial assets and economic resources, slightly less than half of them also come with export restrictions on a set of specific products, largely tied to military use such as nuclear technology, chemicals, or military equipment. This feature allows us to explore whether any collateral damage is due to financial sanctions themselves, or the differential effect of these additional features. As we show, this additional dimension of some sanction measures is indeed relevant.

We uncover three main results. First, while we find some evidence of collateral damage, its extent is typically limited to a subset of activities and actors and does not uniformly affect all trade. As an example, financial sanctions reduce German merchandise exports and imports by about one half as much as they reduce financial flows. However, this reduction is entirely due to those financial sanctions that were accompanied by restrictions on German exports. As a result, such a reduction in trade should not be considered collateral damage. Rather, it is consistent with the idea of sanctions being smart: reducing precisely the activity that they target. Second, the primary channel through which financial sanctions affect cross-border flows is the extensive margin, reducing the number of firms and products engaged in cross-border activities. Third, there are weak anticipation effects in the case of trade in goods and services, but much stronger anticipation effects in cross-border flows of financial assets. Results are almost diametrically opposite in the case of post-sanctions effects.

We make a methodological contribution to the literature on the effects of sanctions by examining the choice researchers face in this literature. The basic identification issue of sanctions owes to the time aggregation of data used by researchers, whether they are annual or monthly, which is usually due to the frequency of the data of the outcome variable of interest. Most papers in this literature use annual data and identify sanctions episodes by the year in which sanctions are imposed and all the remaining years while they are in force. A minority of papers, including both our previous efforts (Besedeš et al. 2017, 2021) as well as Crozet and Hinz (2020), Crozet et al. (2021), Bělin and Hanousek (2021), and Miromanova (2021), identify sanctions episodes at the monthly frequency using the month in which they are imposed.⁴ In our case, all three data sources on outcomes of interest are available at a monthly frequency. This allows us to then compare the effect of sanctions at both the monthly and annual level.

The choice of identification at the monthly or annual level is also tied to how granular the data are. Our merchandise trade data are at the product level with some 16,000 different products. Trade in services as well as financial flows are both identified at the declarant level as well as identifying the service or asset that is being traded. As we will show, preserving such a level of detail and estimating the effect at a monthly frequency is computationally challenging and often prohibitive. This results in a fundamental choice researchers face: either aggregate the time dimension from monthly to annual or aggregate across reporting units.

We also contribute to the fast-growing literature on the economic impact of sanctions. Economists' interest in sanctions and their effects waxes and wanes with their use in practice and, perhaps more importantly, with the importance of their target. The most recent uptake in research pertaining to sanctions is due to sanctions imposed on Russia, both in response to its invasion of Crimea in 2014 and the entire Ukraine in 2022, and Iran, in response to its nuclear arms program and support of terrorism. Crozet and Hinz (2020), Miromanova (2019), Gullstrand (2020), and Drott et al. (2022) examine the effects of various sanctions on Russia, while Haidar (2017), Draca et al. (2017), and Felbermayr et al. (2020a) examine the consequences of recent sanctions on Iran. Unlike the literature focusing on a single targeted country, our effort examines the effect of several episodes of EU sanctions for a single sender country, Germany, which is more in line with Hufbauer and Oegg (2003), Caruso (2003), Yang et al. (2004), and Afesorgbor (2018). In a much broader effort Felbermayr et al. (2020b) build a new database of 729 sanctions episodes between 1950 and 2016 and demonstrate the extent

⁴ Haidar (2017) is in between these two sets of papers identifying sanctions at a quarterly frequency. Drott et al. (2022) are an exemption and use daily data.

to which sanctions reduce trade between countries. Dai et al. (2021) use the same data source to examine the timing of the effect of sanctions showing there are significant negative anticipatory effects preceding sanctions as well as negative lagging effects which take eight years to dissipate. Ahn and Ludema (2019, 2020) provide an analysis of the flipside of sanctions imposed on Russia in 2014. Using data on Russian firms they show that firms targeted by sanctions experience losses in operating revenue, asset value, and employees. Similar to our conclusion, though from a different point of view, they conclude that sanctions on Russia were smart as they had a smaller effect on Russia's macroeconomy than oil prices, indicating that sanctions were affecting intended targets without causing widespread collateral damage.

2. Financial Sanctions in the European Union

Besedeš et al. (2021) describe the implementation of financial sanctions in Germany in practice. Member states of the European Union have committed themselves to a Common Foreign and Security Policy. As a result, foreign policy instruments are imposed by the Council of the EU. Financial sanctions became an available instrument for external action to EU authorities in 1994 when the Treaty of Maastricht entered into force. Among other aspects, the treaty introduced the free movement of capital as a Treaty freedom. Article 63 of the Treaty of the Functioning of the European Union (TFEU) prohibits all restrictions on payments and movement of capital between member states and between member states and third countries, while Article 215 of TFEU allows for the interruption or reduction, in part or completely, of economic and financial relations with one or more third countries.

For our purposes, two features of sanction policies in the European Union are particularly noteworthy. First, while the Council acts by unanimity, regulations are directly applicable in all EU member states and binding in their entirety. As a result, there is only limited scope for potential concerns of endogeneity, where the decision to impose restrictive measures is affected by their expected domestic costs. As the current case of multiple rounds of EU sanctions on Russia due to its invasion of Ukraine in 2022 shows, negotiations leading to a consensus on imposition of sanction can be complicated within the EU. However, while this would ordinarily pose a potential problem for us, our sample ends before this set of sanctions are imposed. In previous sanctions imposed by the EU, either by itself or as a member of the UN, there were few disagreements across EU member states we could find evidence of. As a result, we believe that the sanctions we investigate are plausibly exogenous.

Second, the EU adopts, in practice, a wide range of restrictive measures. These measures often target specific activities and also include restrictions on non-financial activities such as trade embargoes and travel bans. The overwhelming majority of such measures, however, directly and/or indirectly affects cross-border financial relations and are, therefore, officially recorded as a financial sanction, the policy instrument of our interest. Embargoes on exports of specific types of goods, for instance, typically involve restrictions on technical assistance, training and financing. Specific individuals are targeted with travel bans which are often accompanied by other restrictive measures, such as the freezing of funds and financial assets. Since slightly less than half of the financial sanctions episodes in our sample include export restrictions measures, in addition to the freezing of assets and economic resources, we use that feature to explore the differential effect of export restrictions being attached to financial sanctions. The measures are also regularly reviewed and frequently adjusted. Besedeš et al. (2017) show that strengthening of sanctions further reduces German cross-border flows, while weakening results in the opposite effect. In this paper, to save space, we generally limit our attention to the distinction of whether a country is sanctioned or not and ignore the intensity of adopted sanction measures.

3. Data

3.1 Data on sanctions and German cross-border activities

Our analysis is based on four datasets, sourced from the Deutsche Bundesbank and Eurostat. The first source of data consists of information on financial sanctions imposed and enforced by Germany and is primarily obtained from the service center ‘Financial Sanctions’ of the Deutsche Bundesbank as in Besedeš et al. (2017). This unit, which is responsible for the implementation of EU regulations on financial sanctions in Germany, provides a compilation of executive orders and disseminates relevant information to interested parties and the wider public. We augment this data with additional information from official European Union sources.⁵ During our sample period, financial sanctions have been newly imposed on 29 countries. Table 1 provides a list of countries along with a brief description of the measures taken. This table has been updated and extended from Besedeš et al. (2017). Interestingly, while all 29 episodes take the form of financial sanctions by freezing assets and economic resources, 13 of these episodes also provide for some restrictions on exporting to target countries, usually

⁵ Common Foreign and Security Policy Decisions and European Union Regulations are published in the Official Journal of the EU; see <http://eur-lex.europa.eu/homepage.html>.

related to goods that could be used for military purposes, such as nuclear technology, chemicals, or military equipment. In our analysis we will take advantage of this difference across sanctions, a difference that will prove to be important. Figure 1 shows the geographic distribution of sanctioned countries differentiating them by whether sanctions are purely of financial nature or whether they also contain export restrictions.

Sanctions are applied instantaneously, such that there is no time lag between the date of announcement of a sanction and its enforcement. In our empirical analysis, with data at monthly frequency, we code sanctions imposed after the middle of the month as being effective from the beginning of the following month. For six target countries, Liberia, Côte d'Ivoire, Uzbekistan, the Comoros, Eritrea, and the Maldives, sanctions have also been lifted again completely during our sample period and are appropriately coded to reflect the removal of sanctions.

We use monthly data on German imports and exports at the product level taken from Eurostat. To be fully consistent across all our data sets, we use data on imports and exports between January 2001 and September 2020. Eurostat's Comext database provides data on detailed product-level imports and exports for all EU member countries. Products are classified according to EU's 8-digit Combined Nomenclature (CN) classification reflecting some 9,500 products. The data reporting thresholds are established by EU legislation and provide for different reporting thresholds depending on whether data reflect extra- or intra-EU trade in goods. Any extra-EU transaction involving more than €1,000 in value or 1,000 kilograms in net mass must be reported. For intra-EU trade, given the volume of transactions and trade between EU member countries, reporting thresholds are higher and member-specific and are designed to minimize the reporting burden imposed on businesses, especially smaller ones. Four different thresholds are used to determine whether businesses must report their intra-EU trade, with the first three based on the annual value of trade and the last one based on a per-transaction basis. The annual trade thresholds are the exemption, simplification, and statistical value thresholds. Under the exemption threshold member countries can exempt businesses from reporting their trade provided that at least 97% of their intra-EU exports by value and 93% of their intra-EU imports (95% until 2013) are covered and reported according to Eurostat (2020). The simplification threshold allows businesses with annual trade above the exemption threshold but below the simplification threshold to report a limited set of data or use a simplified commodity code. Trade reported by these reporting units may cover at most 6% of a member's total trade. The statistical value threshold, which was discontinued in 2014, allowed member states to

collect the statistical value from their largest reporting units whose overall share of total trade may not exceed 70%. In terms of individual transactions, member states are allowed not to report any transaction that is less than €200. For each product, data report the value of trade, the partner country involved, and a measure of quantity (either units or weight, which we do not use).

We source data on trade in services and cross-border capital flows from two confidential micro data sets from the Deutsche Bundesbank. Given the sensitivity of the business information involved, these data are only accessible, often in anonymized form, at the headquarters of the Bundesbank in Frankfurt, Germany. To compile the balance of payments statistics, the Deutsche Bundesbank collects data on trade in services at the firm level at the monthly frequency. Data are made available through the ‘International Trade in Services Statistics’ (SITS) database which records service transactions using the residence principle, between residents and non-residents, which exceed €12,500 or its equivalent in another currency. The database covers almost the entire population of German service exporters and importers⁶; it comprises data in three of the four modes of the General Agreement of Trade in Services, though in an aggregate fashion precluding the ability of conducting a mode-specific analysis.⁷ Services are categorized according to the sixth edition of IMF’s Balance of Payments and International Investment Position Manual (BPM6).

The SITS database provides information on the reporting unit, the value of each transaction, the type of service involved according to the balance of payments classification, country of destination or origin, and sector of the party required to report. A total of twelve different types of services are reflected in the data: product-related, enterprise-related, personal, intellectual property, telecommunications, construction, transport, insurance, travel, private transfers, transactions by the federation, and other. Taxes are included in reported transaction values. While the original data are reported at the firm level, in order to avoid granularity problems, we perform much of our analysis at the country-month-year level.

Data on cross-border financial flows are similarly sourced from the Deutsche Bundesbank’s balance of payments statistics and are obtained from the Deutsche Bundesbank’s ‘Statistics on International Financial and Capital Transactions’ (SIFCT) database which contains detailed information on financial transactions between Germany and the rest of the

⁶ The Bundesbank supplements the data with estimates for transactions that are below the reporting threshold and for some service categories for which the demanded methodology cannot be reported, such as transportation.

⁷ The three modes are cross-border trade (mode 1), consumption abroad (mode 2), and presence of natural persons (mode 4). The missing mode is commercial presence (mode 3).

world. Data are collected at monthly frequency for the purpose of compiling the balance of payments statistics. The data set is complete with all individuals, firms, and financial institutions in Germany required to report international payments over €12,500. The data reflect the reporting unit, the partner country of each transaction as well as the value and the type of asset involved. To better compare our results to trade in goods and services, we only focus on financial transactions involving German investors, since trade in goods and services reflects the behavior of German firms and consumers. Thus, we focus on capital exports, claims of German investors against foreigners, and capital imports, liabilities of German investors against foreigners.

3.2 Descriptive statistics

The basic reporting unit differs across the three data sets on Germany's cross-border flows. Consequently, in view of this difference, we will largely refer to them as declarants or reporting units, rather than firms or products throughout the paper. Data sourced from Eurostat provide information on Germany's merchandise trade with the basic reporting unit reflecting a product classified under EU's 8-digit CN product classification and are reported on a monthly basis for each of Germany's trade partner country. There are 15,008 8-digit CN product codes in our data set which we aggregate into the ten 1-digit SITC codes. The discrepancy between the number of CN codes in our data and the 9,500 cited above is because CN codes are regularly revised with some codes merged to form new codes, some codes eliminated, and new codes added. Our benchmark results are based on all codes and we explore this issue in the appendix. Data obtained from the Deutsche Bundesbank provide information reported by German businesses or declarants on a monthly basis with information on the type of service or financial transaction involved. There are twelve different categories of services which are further divided into 181 different subcategories. In total, there are 130,702 German declarants reporting data on their trade in services. Financial transactions are classified into six different asset categories with 48 specific types of asset transactions and are reported by 35,407 German declarants.

Given the differences in the number of reporting units, be they products in the case of merchandise trade or German firms and enterprises in the case of trade in services and financial transactions, we base much of our analysis on data aggregated to the country-month-year level. This approach equalizes the analyzed reporting unit to be a country, while keeping the monthly frequency allows us to precisely time the effect of each sanction imposed. As discussed below, when analyzing the effect of sanctions on third country relationships, we will use more granular

data available to us, either the firm-country-month-year level in the case of services trade and financial transactions or the product-country-month-year level in the case of merchandise trade. Accordingly, a statistical entry in our data set is defined as a product, a firm-service pair and a firm-asset pair at the country-month-year level for goods trade, services trade, and financial flows, respectively.

The basic descriptive information on the three data sets we use is presented in the three-part Table 2. For each of the data sets we present information on the full sample as well as two sub-samples of observations, one consisting of observations affected by sanctions and the other consisting of observations that were not affected by sanctions. The last column shows the p-value for a t-test of equality of means between the under-sanctions and not-under-sanctions subsamples.

Table 2a presents summary information on trade in goods which is based on 49,787 country-month-year observations. Unsurprisingly, flows involving goods are larger than those involving services, with an average monthly flow in the full sample of 667 million euros, with 297 million euros due to imports and 371 million euros due to exports. The under-sanction sample consists of 3,593 observations or 7.2% of the full sample. Flows under sanctions have an average size of 152 million euros with 64 million euros falling on imports and 89 million euros on exports. Flows not affected by sanctions are larger, averaging 707 million euros, with 315 million falling on imports and 393 on exports. In terms of cross-sample comparisons, the average flow per entry is significantly larger in the not-under-sanctions sample than in the under-sanctions sample at 190,000€ versus 100,000€

Table 2b shows that there are 50,249 country-month-year observations on trade in services with an average value of 205 million euros, of which 101 million euros are due to imports and 104 million euros are accounted for by exports. For each country-month-year pair, there are on average 454 entries and 321 declarants declaring their trade in services. Flows affected by sanctions account for some 7.2% of observations, 3,593 to be precise. As such they account for a small share of total trade, on average 34 million euros, with 15 million euros imported and 18 million euros exported. Services trade not affected by sanctions averages 218 million euros, with 107 million due to imports and 111 due to exports. Given the relatively small share of flows affected by sanctions, the only dimension in which the equality-of-means test provides useful information are those comparing either per-entry or per-declarant values. We have three such measures, average flow per entry, average number of entries per declarant, and average flow per category per declarant. The average flow per entry is virtually identical in

two subsamples at a quarter of million euros. The average number of entries per declarant is also almost identical at 1.24 in the not-under-sanctions subsample and 1.29 for the under-sanction subsample. Lastly, the average flow per declarant is identical in the two subsamples at 40,000€

Finally, Table 2c presents descriptive information on the financial transactions sample which consists of 32,989 country-month-year observations, with 1,927 observations in the under-sanctions subsample comprising a somewhat smaller fraction (5.8%) of the full sample than was the case for trade in goods and services. Financial flows dwarf both goods and services trade flows with a total of 4,015 million euros in the full sample of which 1,956 million euros are due to inflows and 2,059 million euros due to exports. Whereas the average flow per entry on the trade side was well below a million euros, the average financial flow in the full sample is 12 million euros and 5 million euros in the under the sanctions subsample. The average number of entries per declarant is 1.27 in the not-under-sanctions subsample and 1.06 in the under-sanctions subsample. The average flow per category per declarant is larger in the not-under-sanctions subsample at 4.7 million euros compared to 3.5 million euros in the under-sanctions sample.

4. Trade with Sanctioned Countries

4.1 Benchmark Estimation Specification

We begin our empirical analysis by examining the effect of financial sanctions on trade in services and goods with sanctioned targets. We follow Besedeš et al. (2021) and estimate the following gravity equation using the Poisson pseudo-maximum likelihood estimator (PPML) to take into account observations of zero flows across all our specifications:

$$(1) \quad \text{Flow}_{ct} = \exp[\beta \text{Sanctions}_{ct} + (\gamma X_{ct} +) \eta_c + \phi_t] + \epsilon_{ct},$$

where Flow_{ct} is a measure of the flow of interest, German imports or exports of merchandise goods or services, and inflows or outflows of financial assets, with country c at time t . The exponential function on the right-hand side of equation (1) is due to our use of the PPML estimator. The gravity equation has long been used to empirically examine trade in goods and services. It has also been used to examine cross-border financial flows by Okawa and van

Wincoop (2012) and in direct application to the effects of financial sanctions on financial flows by Besedeš et al. (2021).

Sanctions_{ct} is an indicator variable that takes the value of one when financial sanctions are imposed (and is zero otherwise) against country c at time t , and we include country-specific (η_c) and time-specific (ϕ_t) fixed effects. In several specifications, we also include additional control variables (X_{ct}) at the country-time dimension. The coefficient of interest is β , which measures the effect of sanctions on cross-border flows; a negative and significant coefficient indicates that the adoption of sanctions is associated with fewer transactions between German declarants and their foreign counterparts, *ceteris paribus*. We analyze the data at the country-month-year level by aggregating our detailed data to the country level. We do so to reduce the amount of noise and to compare the three types of flows in goods, services, and financial assets, on as similar a basis as possible; as noted before, the data we use are not based on the same reporting unit with services and financial asset data reported at the firm level and goods trade data reported at the product level. In section 4.3 we examine other possible approaches which preserve the reporter-level detail at both the monthly and annual frequency. To help with the interpretation of results, in every table we report the estimated coefficient, the standard error, and for significantly estimated coefficients the implied relative effect.

4.2 A first take on the effect of financial sanctions

Table 3 reports our benchmark results in Panel A. As in Besedeš et al. (2017) and Besedeš et al. (2021), all time-invariant influences on German flows with a country (such as, for instance, the partner's geographic distance from Germany) are accounted for by country fixed effects, while a comprehensive set of time fixed effects captures monthly variations in capital flows common to all partners. As shown, the point estimates of β are consistently negative and statistically significant. Financial sanctions reduce German imports from targeted countries by 26 percent⁸ and exports to targeted countries by 24 percent. The effect on services trade is somewhat stronger with German services imports from target countries reduced by 31 percent and exports to target countries reduced by 33 percent. Thus, it seems this first investigation of the effect of financial sanctions on trade does indicate collateral damage. Before taking a more detailed approach to this question, the last two columns offer a comparison to the effect of financial sanctions on German financial flows. Consistent with the notion of collateral

⁸ The estimated effect is given by $(e^{-0.306} - 1) \cdot 100 = -26.4\%$.

damage being a secondary effect, financial sanctions have a stronger effect on financial flows, reducing inflows of financial assets from targeted countries by 50 percent and outflows by 48 percent.⁹ Thus, comparing the estimated relative effects, financial sanctions have half as large an effect on trade in goods and two-thirds as large an effect on trade in services as they have on flows of financial assets.

Panels B, C, and D of Table 3 provide some robustness to our results. As is common in the gravity equation literature we use country- and time-specific (that is month-year) fixed effects, also aiming to control for multilateral resistance terms. However, the nature of both our data and our question makes this difficult. Since our data varies at the country-month-year level, ideally we should be employing fixed effects at the country-month-year level to deal with multilateral resistances. However, doing so would preclude us from estimating any effect of sanctions since they are identified at the country-month-year level. While an imperfect approach, in Panel B of Table 3 we include in our specification the atheoretical remoteness index calculated as the GDP-weighted average distance of every partner country. Our estimates of the effect of sanctions change little. Our estimating sample decreases in size due to missing GDP data for a number of countries.

An alternative to using atheoretical measures of remoteness is to follow the work of Freeman et al. (2021) who propose a 2-stage approach to estimating the multilateral resistance terms themselves. Jäkel et al. (2022) use this approach as well in their investigation of the effect of sanctions on Danish exporting firms. This procedure relies on using trade and production data from the ITPD-E database (Borchert et al. 2021) to estimate a structural gravity equation that is used to recover the multilateral resistance terms. We then add these to our regressions. Some caveats are in order. Freeman et al. (2021) develop this approach modeling trade in goods and using annual data on goods trade and production using annual data. They neither do so for trade in services nor cross-border financial flows, which would require setting up an appropriate theoretical model giving rise to a structural gravity equation and having data on domestic production of both services and assets. While a worthwhile goal, it is beyond the scope of our paper. Another complication with this approach is that even for our regressions using data on trade in goods at monthly frequency these MRT terms are not a best fit as they are estimated from annual data. Obtaining the appropriate data to estimate MRT on a monthly frequency is

⁹ Both of these estimates are in line with results in Besedeš et al. (2017) who focus on the effect financial sanctions on financial flows over a short time period.

virtually prohibitive as all requisite data are not available. As a result, we use the estimated MRTs from goods data for both services and financial flows.

Results in panel C show that the inclusion of MRT terms affects the estimated effect of sanctions on trade in goods and services somewhat and reduces the effect on financial flows substantially, lowering the difference between the effect of sanctions on goods and services trade as well as financial flows. However, we caution that the large reduction in the results for financial flows especially should be taken with caution as MRT terms are not estimated from financial flow data. The close correspondence between the estimated effects for goods trade across panels A, B, and C gives us confidence that appropriately estimated MRT terms for services trade and financial flows would show similarly small differences. Given this, with the exception of results in Table 4, for the remainder of our analysis we do not use the estimated MRT terms.

In Panel D we examine whether our results suffer from some omitted variable bias. While our country-level fixed effects control for most standard time-invariant gravity variables, there may be some time-varying variables that could be included. An example would be tariffs in the case of trade in goods and services or similar policy variables. While tariff data are relatively easily available for trade in goods, they are less readily available for trade in services and there are no capital controls imposed on German financial flows (other than sanctions). One alternative we can use, available across the three types of cross-border flows we investigate, is data on various agreements Germany has entered into with other countries. We use the Baier and Bergstrand (2007) Database on Economic Integration Agreements¹⁰ to identify agreements affecting Germany's trade in goods. While the database identifies six different types of agreements, we use a single dummy to identify the existence of any trade agreement between Germany and a trading partner. For trade in services we use information on services commitments in regional trade agreements made available by the WTO.¹¹ Finally, to identify similar agreements affecting financial flows we use UNCTAD's International Investment Agreements Navigator data¹² to identify bilateral investment treaties (BIT) and treaties with investment provisions (TIP) Germany has entered into. Across all our regressions we also include the GDP of the partner country. As seen from our results in Panel D of Table 3, the

¹⁰ <https://sites.nd.edu/jeffrey-bergstrand/database-on-economic-integration-agreements/>

¹¹ https://www.wto.org/english/tratop_e/serv_e/dataset_e/dataset_e.htm

¹² <https://investmentpolicy.unctad.org/international-investment-agreements>

inclusion of these bilateral time-varying macroeconomic variables affects our estimates of the effect of sanctions very little, usually marginally increasing their magnitude.

In Table 4, we explore the effect of different identification choices we can make in order to identify the effect of sanctions. In these specifications, we aim to include more restrictive pairwise product/declarant-time fixed effects and product/declarant-country fixed effects as controls to account for unobserved heterogeneity.¹³ Panel A contains the estimates from Panel C of Table 3: the effect of sanctions estimated with data aggregated to the country-month-year level. In Panel B of Table 4, we preserve the reporting unit disaggregation (either product or declarant) of our raw data, but aggregate the time dimension to annual frequency, as is more commonly done in the literature. This increases the number of observations in our sample considerably. There are now 7.5 and 15 million observations in our imports and exports of goods regressions, about half as much in corresponding trade in services regressions, and around 830,000 and 920,000 observations in financial flow regressions. The increase in the number of observations results in still significantly estimated effects for inflows and outflows of financial assets, while the effect is only significant for exports of goods. Given that we are estimating a structural gravity equation, the reduction in significance may be due to the large number of zeros that are introduced in the dataset or the noisier identification of financial sanctions.

The biggest challenge for identification is estimating the effect of sanctions on the most disaggregated data available to us, preserving both the reporting unit disaggregation as well as the time disaggregation at the monthly frequency, if the researcher generates a balanced panel, a virtual necessity when estimating the structural gravity model. We can only do so for imports and exports of goods which now have 89 and 195 million observations used in estimation. For goods trade we obtain similar results. The effect on imports is not significant, while that for exports is and of similar magnitude with data aggregated to annual frequency. We cannot obtain the results at this level of disaggregation for trade in services and financial flows as they are computationally prohibitive. Since we use confidential data, we have to use computing power provided by Bundesbank's Research Center which is unable to process our datasets which are

¹³ These more restrictive fixed effects come with a computational advantage as a large proportion of observations in the more disaggregated data are dropped as they are either singletons or separated by a fixed effect. In the most disaggregated data at the product-country-month-year level 88% of observations are dropped in our imports regression and 75% in the exports regression. With the less restrictive set of fixed effect (country, product, and month-year) fewer observations are dropped, 25% for imports and 21 % for exports. The estimated effect of sanctions is similar in the exports regression, while that for imports is twice as large and statistically highly significant.

huge. Imports of services dataset contains 8.5 billion observations while that of exports of services almost 25 billion observations in a balanced panel required for structural gravity. Financial inflows are slightly under 2 billion observations, while outflows are somewhat over 3 billion observations.¹⁴

Given this discussion we are left with a choice: either to aggregate the reporter dimension preserving the monthly frequency of data allowing for a more precise identification of sanctions or to preserve the reporter disaggregation while aggregating to annual frequency. The latter option adds more noise to the data, irrespective of the time frequency, as there are likely to be many reporters (firms, products, entities) with zero flows rather than positive ones. The former option offers more value as monthly frequency allows for a more precise identification of the effect of sanctions. The main concern is the partial-year effect similar to that discussed by Bernard et al. (2017). At the annual frequency sanctions which began in January would be treated similarly as those which started in December. However, the former would have affected flows for an entire year, while the latter would have been in place for only one month but treated as if they were in place for the entire year. In addition, since much of the literature uses annual data, there is greater value in examining the effects at the monthly frequency. An additional choice we have to make is how to incorporate MRT terms. Panel C of Table 3 shows that incorporating MRT terms explicitly does not significantly alter estimates for goods trade regressions. This is important as MRT terms are estimated from goods trade data. Given that we cannot estimate MRT terms separately for services and cross-border financial flows and that MRTs only vary on annual frequency, we prefer to rely on the fixed effects we can include as the main specification. As a result of both of these choices, we focus the remainder of our analysis at the country-month-year level and consider panel A of Table 3 to be our benchmark results.

4.3 Intensive and extensive margins

In Table 5 we expand our analysis to cover other quantitative features of Germany's bilateral cross-border relationships. We decompose aggregate flows with a partner country into different margins that we summarized in Table 2. Understanding how sanctions affect the various margins can help us identify the channels through which sanctions reduce cross-border flows. The margins we investigate are the intensive margin, or the average value per entry, and

¹⁴ Another computational challenge is our use of STATA which is limited to 2.1 billion observations.

the extensive margin, which we define as the number of entries. Since both the margins are summaries of disaggregated data, we use our reporting unit data at the monthly frequency to calculate them. Specifically, we use data at the product/declarant-month-year level to calculate the intensive margin, the average flow per entry, at the country-month-year level, and we do the same for the extensive margin by calculating the number of entries within each country-month-year cell. We calculate two additional extensive margins, counting the number of declarants or products (depending on the flow) and the number of categories (for service trade), industries (goods trade), or asset classes (for cross-border financial flows). We present evidence on more measures of the extensive margin as sanctions affect that margin more so than the intensive margin.

We estimate variants of equation (1), but with each different margin serving as the dependent variable. The two regressions involving the value of trade and reflecting the intensive margin, in rows one and three, are estimated using PPML. The remaining three results, in rows two, four, and five, contain results for various count measures. As Santos Silva et al. (2014) point out, such measures have both a lower and an upper bound with resulting partial effects of explanatory variables on the conditional mean of the dependent variable not being constant. We use the Flex estimator suggested by Santos Silva et al. (2014) to deal with this issue.

To save space, we only report estimates for the coefficient of interest, β . The dependent variable is tabulated in the first column on the left of the table. The first row reproduces, for comparison, estimates from Table 3, showing the effect of sanctions on total cross-border flows at the country-month-year level. The remaining rows show the estimated effect of financial sanctions on various margins. To be precise, the second row presents the estimated effect of sanctions on the number of entries underlying every country-month-year observation. In the case of merchandise trade this is the number of product observations for each country-month-year pair. In the case of trade in services this is the number of firm-service observations for each country-month-year pair and in the case of financial flows it is the number of firm-asset observations for each country-month-year pair.

Among the various margins, financial sanctions have the most consistent (negative) effect on the number of entries and the number of declarants or products. This holds for all types of cross-border flows we examine. Exports of goods experience a 30 percent decline in the number of entries (products), while the effect on imports is not statistically distinguishable from zero. The effect on services trade is somewhat weaker, with the number of entries for imports decreasing by 22 percent and exports by about 24 percent. The number of entries for

financial-asset inflows and outflows decreases by about 30 percent. The relative effects of sanctions on the number of declarants, products in the case of merchandise trade and firms in the case of services and financial asset transactions, are almost identical in terms of magnitude and significance to the effects on the number of entries.¹⁵ Sanctions reduce the number of asset classes by more than 60 percent. They also reduce the number of imported service categories by 42 percent, while having no significant impact on exports of service categories, or industries in the case of trade in goods. On the intensive margin, sanctions only have an effect on exports of goods, reducing average exports by about 27 percent. Thus, the main channel through which financial sanctions affect German cross-border flows is through the extensive margin, reducing the number of declarants that engage in cross-border flows while sanctions are in place. Only exports of goods seem to be affected on the intensive margin.

4.4 Collateral damage or smart sanctions?

Our main question of interest is whether financial sanctions produce spillover effects (or collateral damage) on trade in goods and services. If there are no effects, one could make a strong case that financial sanctions are indeed smart in the sense that their economic effect and damage is narrowly focused on reducing cross-border financial flows. Our results so far imply that financial sanctions do create collateral damage. They reduce both imports and exports of goods and services, having half as large an effect on merchandise trade and two-thirds as large an effect on services trade as they do on financial flows.

However, this may not be a complete picture. The reason is that, as noted above, 13 of the 29 sanctions episodes in our data set also have provisions that restrict trade in certain types of goods, usually related to their military use. It is entirely possible that the negative effect of financial sanctions we have identified on trade in goods and services is due to these episodes that have non-finance related stipulations. In order to investigate this possibility in more detail, we estimate the following specification:

$$(2) \quad \text{Flow}_{ct} = \exp[\beta \text{Sanctions}_{ct} + \gamma \text{SanctionsExports}_{ct} + \eta_c + \phi_t] + \epsilon_{ct},$$

¹⁵ In the case of trade in goods, the results for the number of entries and the number of declarants are identical as the two datasets are identical. The declarant in the goods trade dataset is the product.

where we add a separate dummy variable, $\text{SanctionsExports}_{ct}$, which identifies the 13 sanctions episodes with export restrictions. The estimate of the coefficient γ is then the differential effect those additional restrictions have on trade in goods and services.

Results shown in Table 6 indicate that allowing for this differential effect bolsters the argument that financial sanctions affecting Germany's cross-border flows in trade and finance are smart and cause only moderate collateral damage. The strongest results are obtained for merchandise trade. Financial sanctions do not affect either imports or exports of goods unless they are accompanied by specific export restrictions. In such cases, these additional restrictions reduce imports by 32 percent and exports by 26 percent. As far as trade in services is concerned, imports of services are no longer affected in a statistically meaningful way. Financial sanctions do reduce Germany's exports of services with no differential effect of sanctions that entail additional export restrictions. The effect of financial sanctions on financial flows increases in this expanded specification to 56 percent and 54 percent for inflows and outflows, with no significant differential effect of sanctions with export restrictions.

We make several observations about the potential for collateral damage caused by financial sanctions. The strongest evidence we have for them is in the case of service exports which are reduced by a third by any kind of financial sanctions. Given that some services are needed for the conduct of cross-border financial flows, such as financial services, this effect is perhaps not surprising. Note that in the case of merchandise trade, the only collateral damage we can identify is in the case of sanctions that also have export restrictions attached to them. Given that these additional restrictions are restrictions on exports, it seems difficult to argue that financial sanctions create collateral damage. The only remaining case of collateral damage is then due to the result that export restrictions reduce imports of goods. Our conclusion is that the extent of collateral damage caused by financial sanctions is limited and that these sanctions are likely smart: their main and strongest effects are to reduce their primary target – financial flows.¹⁶

¹⁶ Goods and services trade is at least to some extent connected as some firms likely engage in both, especially those firm trading products that require maintenance. Unfortunately, our data prevent us from being able to examine the effect of sanctions on such firms as the two datasets, on trade in goods and trade in services, are separate and with different reporting units.

4.4 Pre- and post-sanction effects

While the bulk of the effect of sanctions is contemporaneous, it is possible that they may be preceded by anticipation effects and followed by lingering effects affecting cross-border flows after they are removed. In fact, Dai et al. (2021) find significant pre- and post-sanctions effects in the case of trade sanctions, with the post-sanctions effects lingering for 8 years. They argue that including pre- and post-effects increases the estimates of the contemporaneous effect. In our earlier work on the effects of German financial sanctions we found evidence of anticipatory effects reducing cross-border financial flows, but also large positive effects of increases in cross-border flows after sanctions were removed (Besedeš et al. 2017).

We now investigate the possibility of pre- and post-sanctions effects on cross-border flows of goods, services, and capital. We do so by adding two new variables to equation (1): a pre-sanctions dummy and a dummy for the removal of sanctions.. The pre-sanctions period dummy variable identifies the period between the event that was identified as the trigger event that led to the imposition of sanctions, such as the annexation of Crimea in the case of sanctions imposed on Russia, and the imposition of sanctions. This dummy identifies the period from the point when the specter of sanctions first arose and their eventual imposition. For 13 of the 29 sanctions episodes in our dataset there is a period between the two events varying from virtually instantaneous imposition in the case of Russia and the Comoros to 12 months in the case of Lebanon and Eritrea, with an average of 3.6 months and a median of 2 months. For the other 16 episodes, there are no possible anticipation effects since there is no specific event that is cited in the announcement of sanctions. Of the 29 episodes of sanctions imposed by Germany during the period we examine, only six were removed, on the Comoros, Côte d'Ivoire, Eritrea, Liberia, the Maldives, and Uzbekistan. For these six episodes we add a dummy identifying the post-sanctions period, when they were no longer in place.

Our results in Table 7 indicate almost no change in the contemporaneous effect of sanctions when we include the anticipation dummy and a sanctions removal dummy. There are significant anticipation effects for exports of goods (reduction of 22 percent) and services (reduction of 26 percent), and financial inflows (reduction of 52 percent).¹⁷ Note that these anticipation effects are similar to contemporaneous effects for exports of goods and financial inflows, and somewhat weaker in the case of exports of services. Curiously, we find evidence of additional negative post-sanctions effects in the case of imports of goods and services. Goods

¹⁷ An alternative explanation of the anticipation effect is that the trigger event itself increases country risk causing a reduction in flows even if no sanctions are forthcoming.

imports decrease by an additional 50 percent and service imports decrease by an additional 74 percent *after* the removal of sanctions. Both lingering effects of sanctions are much larger than the contemporaneous effects. In the case of financial flows, we find no evidence of post-sanctions effects of any kind. In no case have we found a significant positive effect of the removal of sanctions. A possible explanation could be that it is easier to restart financial relationships after sanctions are over, while those involving goods or services are more difficult as they may involve higher fixed costs. The asymmetry between imports and exports for goods and services may be due to the fact that it may be easier for German firms to pay such costs to restart those relationships than is the case for firms from these six countries, all of which are low-income developing countries.

We can draw two conclusions from our investigations of pre- and post-sanctions effects. First, the anticipation effects of financial sanctions on trade in goods and services only affect exports, while the post-sanctions effects are stronger for imports of both goods and services. In this dimension, our results indicate weaker effects than those found by Dai et al. (2021), though they are identifying first order effects of trade sanctions on trade, while we are identifying second order effects of financial sanctions on trade, or as we put it earlier, collateral damage of financial sanctions. Second, anticipation effects are much stronger for the effect of financial sanctions on inflows of capital, the primary target of financial sanctions, while there are no post-sanction effects on either financial inflows or outflows.

5. Conclusions

A recent increase in the use of sanctions as a foreign policy tool has increased researchers' attention to their effects. While much effort has gone into uncovering the direct effects of sanctions on the activity primarily targeted by them, less attention has been paid to the extent the effects of sanctions spill over into other activities. We examine the extent of collateral damage in trade in goods and services resulting from German financial sanctions. The aim of financial sanctions is to restrict the cross-border flow of financial activities. Such sanctions could create collateral damage by reducing trade in goods and services. A simple link could be that the presence of financial sanctions increases the risk of doing business, any business, with the sanctioned country resulting in a broad reduction in economic interaction between the sender of sanctions and its target. We find limited evidence of such collateral damage effects.

Financial sanctions reduce the cross-border capital flows by some 50 percent and seemingly cause collateral damage by having half the effect on trade in goods and two-thirds as large a negative effect on trade in services. But the collateral damage is almost entirely due to sanctions episodes where financial sanctions are accompanied by export restrictions. Since export restrictions are designed to limit trade, one can hardly think of these effects as being evidence of collateral damage.

The primary channel through which financial sanctions affect cross-border flows is the extensive margin reducing the number of firms or products engaged in cross-border flows when sanctions are in effect. We find weak anticipation effects and stronger lagging negative effects for imports of both goods and services. Anticipation effects are much stronger when it comes to financial assets, but there are no lagging post-sanctions effects on financial assets.

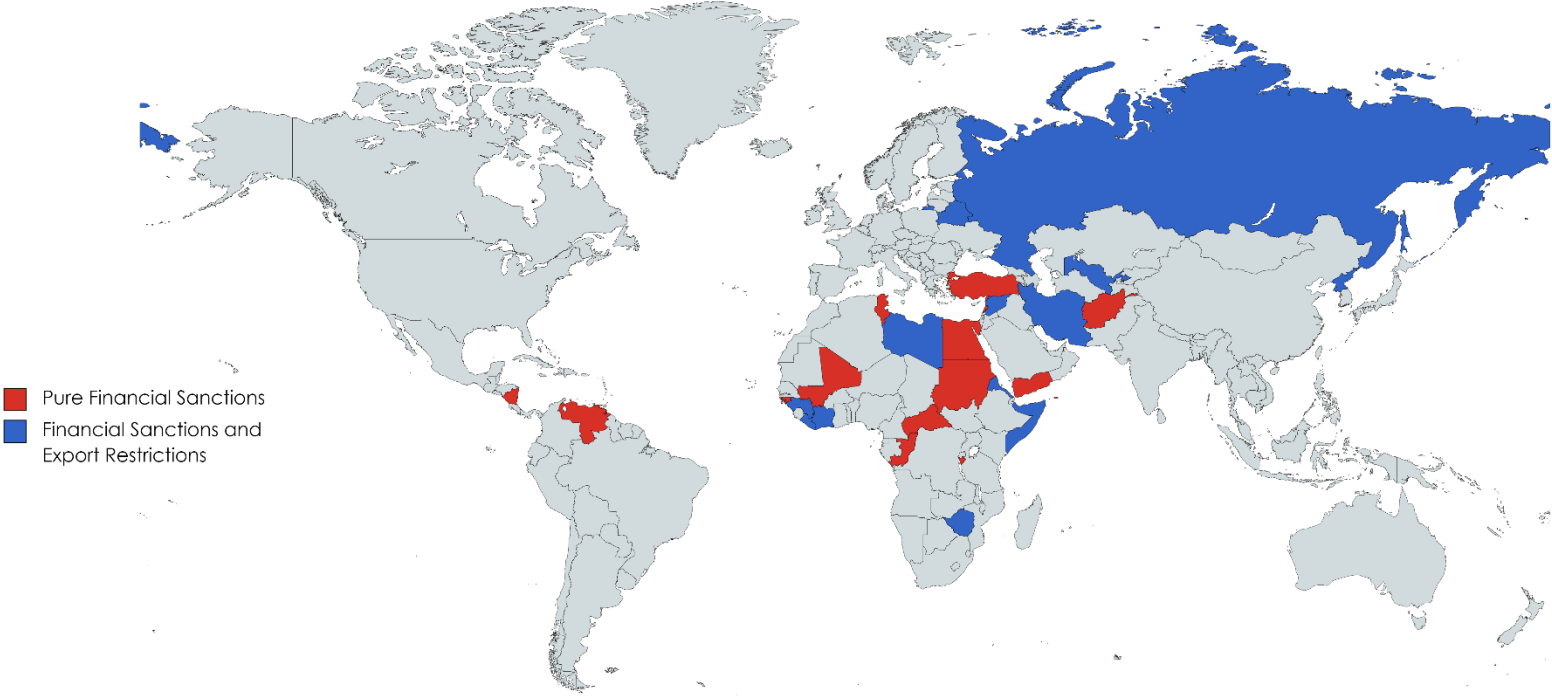
Our aim in this paper was to investigate whether financial sanctions cause collateral damage or can be thought of as smart sanctions if there is no collateral damage. We conclude that there is limited evidence of collateral damage, but it is not widespread and is contained to certain specific services and goods. As such, on the whole, financial sanctions imposed by Germany do seem to be smart with their effects mostly concentrated on the targeted activity.

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Figure 1: Geographic Distribution of Sanctions



Created with mapchart.net

Table 1: List of Financial Sanctions, 2001-2020

| Country | First announcement (Lifted) | Measures taken | Cause cited in declaration | Date of event | Sanctions initially imposed by |
|------------------|--|--|---|----------------------|---------------------------------------|
| Somalia | 27 January 2003 | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment | Situation in Somalia | | UN |
| Liberia | 4 September 2003 (20 June 2016) | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment | Situation in Liberia | | UN |
| Congo, Dem. Rep. | 29 September 2003 | Freezing of assets and economic resources of natural persons and establishments | Violation arms embargo | | UN |
| Sudan | 26 January 2004 | Freezing of assets and economic resources of natural persons | Situation in Sudan | | UN |
| Zimbabwe | 19 February 2004 | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment | Situation in Zimbabwe | | EU |
| Côte d'Ivoire | 31 January 2005 (9 June 2016) | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment | Ceasefire violation | 15 November 2004 | UN |
| Uzbekistan | 14 November 2005 (15 December 2009) | Freezing of assets and economic resources; export restriction on goods related to nuclear technology | Massacre in Andijan | 13 May 2005 | EU |
| Lebanon | 21 February 2006 | Freezing of assets and economic resources | Assassination of former Lebanese Prime Minister | 14 February 2005 | UN |
| Belarus | 18 May 2006 | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment | Presidential elections | 19 March 2006 | EU |
| Iran | 2 February 2007 | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment, chemicals and other resources (gold, silver, ...) | Deterioration of human rights situation in Iran | | UN |
| Korea, Dem. Rep. | 27 March 2007 | Freezing of assets and economic resources of natural persons and establishments; export restriction on luxury goods and goods related to nuclear technology | Nuclear test | 9 October 2006 | UN |

| | | | | | |
|--------------------------|------------------------------------|--|--|-------------------|----|
| Comoros | 17 March 2008 (24 July 2008) | Freezing of assets and economic resources of natural persons | Invasion of Anjouan | 17 March 2008 | EU |
| Guinea | 22 December 2009 | Freezing of assets and economic resources of natural persons; export restriction on military equipment | Violent repression | 28 September 2009 | EU |
| Eritrea | 26 July 2010 (10 December 2018) | Freezing of assets and economic resources; export restriction on military equipment | Border dispute with Djibouti and support of Somalia; AU summit declaration calling for sanctions | 3 July 2009 | UN |
| Tunisia | 4 February 2011 | Freezing of assets and economic resources of natural persons | Situation in Tunisia | 18 December 2010 | EU |
| Libya | 2 March 2011 | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment | Situation in Libya | 17 February 2011 | UN |
| Egypt | 21 March 2011 | Freezing of assets and economic resources of natural persons | Situation in Egypt | 25 January 2011 | EU |
| Syria | 9 May 2011 | Freezing of assets and economic resources of natural persons and establishments; export restriction on military equipment, chemicals and other resources (gold, silver, ...) | Repression of the civilian population | | EU |
| Afghanistan | 1 August 2011 | Freezing of assets and economic resources of natural persons and establishments | Situation in Afghanistan | 17 June 2011 | UN |
| Guinea-Bissau | 3 May 2012 | Freezing of assets and economic resources of natural persons | Coup d'etat | 12 April 2012 | EU |
| Russia | 5 March 2014 | Freezing of assets and economic resources of natural persons and establishments; export restriction on oil drilling machinery, chemicals and other natural resources | Annexation Crimea | 3 March 2014 | EU |
| Central African Republic | 10 March 2014 | Freezing of assets and economic resources of natural persons and establishments | Situation in the Central African Republic | 5 December 2013 | UN |
| Yemen | 18 December 2014 | Freezing of assets and economic resources of natural persons | Political situation in Yemen | | UN |
| Burundi | 1 October 2015 | Freezing of assets and economic resources of natural persons | Violent repression | | EU |
| Mali | 28 September 2017 | Freezing of assets and economic resources of natural persons | Violent repression | | UN |
| Venezuela | 13 November 2017 | Freezing of assets and economic resources of natural persons | Violation of democracy, justice and human rights | | EU |

| | | | | | |
|-----------|--------------------------------|--|----------------------------|--|----|
| Maldives | 16 July 2018 (17 June 2019) | Freezing of assets and economic resources of natural persons | Violation of human rights | | EU |
| Nicaragua | 14 October 2019 | Freezing of assets and economic resources of natural persons | Violent repression | | EU |
| Turkey | 11 November 2019 | Freezing of assets and economic resources of natural persons | Oil drilling in open water | | EU |

Source: Deutsche Bundesbank, Service center 'Financial Sanctions' and own investigations

Table 2a: Descriptive Statistics Trade in Goods

| | Full Sample | | | Not Sanctioned | | | Under Sanction | | | t-test (p-value) |
|---|-------------|--------|--------------|----------------|--------|--------------|----------------|--------|--------------|---------------------|
| | Obs. | Mean | Std. Dev. | Obs. | Mean | Std. Dev. | Obs. | Mean | Std. Dev. | |
| Total Flows (Mn. €) | 49,787 | 667.35 | 2019.5 | 46,194 | 707.40 | 2085.5 | 3,593 | 152.42 | 559.15 | 0.00 |
| Avg. Flow per Entry (Mn. €) | 49,787 | 0.18 | 1.00 | 46,194 | 0.19 | 1.04 | 3,593 | 0.10 | 0.22 | 0.00 |
| Products (Number) | 49,787 | 1492.7 | 1964.3 | 46,194 | 1545.9 | 2007.8 | 3,593 | 809.55 | 1064.6 | 0.00 |
| Industries (Number) | 49,787 | 7.89 | 2.38 | 46,194 | 7.87 | 2.42 | 3,593 | 8.06 | 1.80 | 0.00 |
| | | | | | | | | | | |
| Imports (Mn. €) | 49,787 | 296.49 | 954.03 | 46,194 | 314.58 | 985.58 | 3,593 | 63.90 | 255.24 | 0.00 |
| Exports (Mn. €) | 49,787 | 370.86 | 1116.5 | 46,194 | 392.82 | 1152.9 | 3,593 | 88.52 | 310.92 | 0.00 |
| | | | | | | | | | | |
| Categories (Mn. €) | | | | | | | | | | |
| – Food and live animals | 44,019 | 41.04 | 151.28 | 40,744 | 43.79 | 156.85 | 3,275 | 6.83 | 160.57 | 0.00 |
| – Beverages and tobacco | 37,890 | 7.18 | 22.37 | 35,150 | 7.65 | 23.15 | 2,740 | 1.16 | 2.86 | 0.00 |
| – Crude materials, inedible, except fuels | 39,875 | 21.93 | 70.86 | 36,642 | 23.55 | 73.65 | 3,233 | 3.63 | 9.01 | 0.00 |
| – Mineral fuels, lubricants and related materials | 31,722 | 47.61 | 191.06 | 29,495 | 46.03 | 185.25 | 2,227 | 68.57 | 254.97 | 0.00 |
| – Animal and vegetable oils, fats and waxes | 26,229 | 3.48 | 13.02 | 24,561 | 3.69 | 13.42 | 1,668 | 0.38 | 1.27 | 0.00 |
| – Chemicals and related products, n.e.s. | 46,072 | 104.70 | 366.39 | 42,543 | 111.92 | 379.89 | 3,529 | 17.74 | 67.82 | 0.00 |
| – Manufactured goods | 46,860 | 97.80 | 292.80 | 43,362 | 104.36 | 302.88 | 3,498 | 16.48 | 64.42 | 0.00 |
| – Machinery and transport equipment | 48,487 | 312.84 | 950.28 | 44,941 | 333.29 | 982.83 | 3,546 | 53.60 | 182.14 | 0.00 |
| – Miscellaneous manufactured articles | 47,811 | 82.73 | 253.62 | 44,241 | 88.37 | 262.56 | 3,570 | 12.83 | 42.74 | 0.00 |
| – Commodities and transactions, n.e.s. | 23,662 | 5.91 | 36.63 | 21,990 | 6.32 | 37.93 | 1,672 | 0.56 | 6.32 | 0.00 |

Notes: The unit of observation is a country-month-year pair. If not noted otherwise, values refer to the sum of exports and imports.

Table 2b: Descriptive Statistics Trade in Services

| | Full Sample | | | Not Sanctioned | | | Under Sanction | | | t-test (p-value) |
|---|-------------|--------|-----------|----------------|--------|-----------|----------------|--------|-----------|---------------------|
| | Obs. | Mean | Std. Dev. | Obs. | Mean | Std. Dev. | Obs. | Mean | Std. Dev. | |
| Total Flows (Mn. €) | 50,249 | 204.58 | 843.67 | 46,656 | 217.79 | 873.56 | 3,593 | 33.76 | 117.29 | 0.00 |
| Entries (Number) | 50,249 | 453.47 | 1142.6 | 46,656 | 479.14 | 1179.38 | 3,593 | 120.05 | 277.82 | 0.00 |
| Avg. Flow per Entry (Mn. €) | 50,249 | 0.24 | 0.78 | 46,656 | 0.23 | 0.64 | 3,593 | 0.25 | 1.77 | 0.33 |
| Declarants (Number) | 50,249 | 320.96 | 752.81 | 46,656 | 338.74 | 776.40 | 3,593 | 90.10 | 202.24 | 0.00 |
| Avg. Number of Entries per Declarant | 50,249 | 1.24 | 0.18 | 46,656 | 1.24 | 0.18 | 3,593 | 1.29 | 0.21 | 0.00 |
| Categories (Number) | 50,249 | 8.97 | 3.41 | 46,656 | 8.98 | 3.45 | 3,593 | 8.84 | 2.88 | 0.02 |
| Avg. Flow per Category per Declarant (Mn. €) | 50,249 | 0.04 | 0.17 | 46,656 | 0.04 | 0.15 | 3,593 | 0.04 | 0.28 | 0.13 |
| | | | | | | | | | | |
| Imports (Mn. €) | 50,249 | 100.48 | 431.00 | 46,656 | 107.07 | 446.11 | 3,593 | 14.87 | 76.07 | 0.00 |
| Exports (Mn. €) | 50,249 | 104.10 | 436.53 | 46,656 | 110.71 | 452.02 | 3,593 | 18.20 | 62.44 | 0.00 |
| | | | | | | | | | | |
| Categories (Mn. €) | | | | | | | | | | |
| – Product-related | 39,346 | 27.53 | 121.67 | 36,571 | 29.16 | 124.90 | 2,775 | 5.96 | 61.97 | 0.00 |
| – Enterprise-related | 42,456 | 38.82 | 185.24 | 39,292 | 41.65 | 192.24 | 3,164 | 3.64 | 13.33 | 0.00 |
| – Personal | 29,447 | 4.06 | 18.46 | 27,319 | 4.31 | 18.76 | 2,128 | 0.91 | 1.37 | 0.00 |
| – Intellectual property | 33,121 | 22.93 | 120.55 | 30,804 | 24.48 | 124.83 | 2,317 | 2.32 | 10.26 | 0.00 |
| – Telecommunications | 36,798 | 18.33 | 69.15 | 34,417 | 19.46 | 71.34 | 2,381 | 1.98 | 7.98 | 0.00 |
| – Construction | 34,897 | 11.15 | 97.45 | 32,184 | 11.65 | 101.18 | 2,713 | 5.16 | 26.90 | 0.00 |
| – Transport | 44,808 | 45.37 | 144.68 | 41,641 | 48.21 | 149.54 | 3,167 | 8.08 | 25.26 | 0.00 |
| – Insurance | 35,505 | 55.53 | 355.03 | 33,553 | 58.49 | 364.91 | 1,952 | 4.75 | 32.11 | 0.00 |
| – Travel | 38,231 | 4.74 | 18.55 | 35,867 | 4.91 | 19.03 | 2,364 | 2.22 | 8.05 | 0.00 |
| – Private transfers | 38,533 | 3.84 | 16.97 | 35,108 | 4.00 | 16.25 | 3,425 | 2.15 | 22.97 | 0.00 |
| – Transactions by the federation | 47,227 | 18.21 | 82.01 | 43,669 | 19.34 | 84.62 | 3,558 | 4.43 | 34.36 | 0.00 |
| – Other | 30,268 | 13.62 | 48.84 | 28,462 | 14.43 | 50.26 | 1,806 | 0.84 | 2.37 | 0.00 |

Notes: The unit of observation is a country-month-year pair. If not noted otherwise, values refer to the sum of exports and imports.

Table 2c: Descriptive Statistics Financial Flows

| | Full Sample | | | Not Sanctioned | | | Under Sanction | | | t-test (p-value) |
|---|-------------|--------|-----------|----------------|--------|-----------|----------------|--------|-----------|---------------------|
| | Obs. | Mean | Std. Dev. | Obs. | Mean | Std. Dev. | Obs. | Mean | Std. Dev. | |
| Total Flows (Mn. €) | 32,989 | 4014.9 | 14,900 | 31,062 | 4258.3 | 15,400 | 1,927 | 91.32 | 352.39 | 0.00 |
| Entries (Number) | 32,989 | 95.14 | 190.35 | 31,062 | 100.06 | 194.93 | 1,927 | 15.71 | 32.62 | 0.00 |
| Avg. Flow per Entry (Mn. €) | 32,989 | 11.87 | 38.69 | 31,062 | 12.33 | 37.99 | 1,927 | 4.58 | 48.09 | 0.00 |
| Declarants (Number) | 32,989 | 59.37 | 108.92 | 31,062 | 62.24 | 111.44 | 1,927 | 13.11 | 25.16 | 0.00 |
| Avg. Number of Entries per Declarant | 32,989 | 1.26 | 0.30 | 31,062 | 1.27 | 0.31 | 1,927 | 1.06 | 0.14 | 0.00 |
| Asset Classes (Number) | 32,989 | 3.21 | 1.75 | 31,062 | 3.28 | 1.75 | 1,927 | 2.13 | 1.17 | 0.00 |
| Avg. Flow per Category per Declarant (Mn. €) | 32,989 | 4.64 | 27.27 | 31,062 | 4.71 | 25.48 | 1,927 | 3.53 | 4.63 | 0.07 |
| | | | | | | | | | | |
| Inflows (Mn. €) | 32,989 | 1956.2 | 7325.8 | 31,062 | 2074.8 | 7533.5 | 1,927 | 45.16 | 185.48 | 0.00 |
| Outflows (Mn. €) | 32,989 | 2058.7 | 7651.3 | 31,062 | 2184.6 | 7868.0 | 1,927 | 46.16 | 175.95 | 0.00 |
| | | | | | | | | | | |
| Asset Classes (Mn. €) | | | | | | | | | | |
| – Bonds | 24,859 | 2620.9 | 9071.9 | 23,730 | 2742.2 | 9267.7 | 1,129 | 71.01 | 197.05 | 0.00 |
| – Money market instruments | 8,075 | 1938.0 | 4257.8 | 8,000 | 1956.0 | 4273.7 | 75 | 16.30 | 29.01 | 0.00 |
| – Equity | 18,806 | 1552.3 | 6167.8 | 18,229 | 1598.2 | 6261.0 | 577 | 100.92 | 273.45 | 0.00 |
| – Collective investment | 7,391 | 1941.5 | 7443.1 | 7,378 | 1945.0 | 7449.2 | 13 | 0.49 | 1.15 | 0.35 |
| – Foreign direct investment | 22,548 | 284.11 | 1457.2 | 21,540 | 296.30 | 1489.7 | 1,008 | 23.64 | 108.11 | 0.00 |
| – Other | 24,314 | 69.90 | 488.55 | 23,012 | 73.31 | 501.72 | 1,302 | 9.61 | 64.78 | 0.00 |

Notes: The unit of observation is a country-month-year pair. If not noted otherwise, values refer to the sum of inflows and outflows.

Table 3: The Effect of (Financial) Sanctions on Trade in Goods/Services/Financial Flows

| | Goods | | Services | | Financial Flows | |
|---|--------------------------------|--------------------------------|-------------------------------|--------------------------------|--------------------------------|--------------------------------|
| | Imports | Exports | Imports | Exports | Inflows | Outflows |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Panel A: Benchmark | | | | | | |
| (Financial) Sanctions | -0.306*** (0.077) -26.4% | -0.269*** (0.072) -23.6% | -0.374* (0.192) -31.2% | -0.401*** (0.084) -33.0% | -0.694*** (0.157) -50.0% | -0.645*** (0.163) -47.5% |
| Observations | 51,192 | 51,192 | 52,140 | 52,140 | 48,822 | 48,822 |
| Pseudo R² | 0.987 | 0.989 | 0.967 | 0.963 | 0.972 | 0.972 |
| Panel B: Remoteness | | | | | | |
| (Financial) Sanctions | -0.300*** (0.075) -25.9% | -0.345*** (0.062) -29.2% | -0.397** (0.179) -32.8% | -0.401*** (0.082) -33.0% | -0.634*** (0.147) -47.0% | -0.595*** (0.143) -44.6% |
| Remoteness | -0.014* (0.008) | -0.037*** (0.010) | -0.017 (0.011) | -0.013 (0.009) | -0.027 (0.027) | -0.031 (0.025) |
| Observations | 40,812 | 40,812 | 40,812 | 40,812 | 40,128 | 40,128 |
| Pseudo R² | 0.987 | 0.991 | 0.965 | 0.961 | 0.973 | 0.973 |
| Panel C: Multilateral Resistance Terms | | | | | | |
| (Financial) Sanctions | -0.281*** (0.085) -24.5% | -0.288*** (0.063) -25.0% | -0.299** (0.142) -25.8% | -0.373*** (0.085) -31.1% | -0.447*** (0.148) -36.0% | -0.512*** (0.141) -40.1% |
| Multilateral Resistance Terms | -0.001* (0.000) | 0.017** (0.008) | 0.001*** (0.000) | 0.018 (0.014) | 0.001** (0.000) | 0.114* (0.058) |
| Observations | 38,664 | 40,944 | 38,664 | 40,944 | 38,016 | 40,260 |
| Pseudo R² | 0.987 | 0.989 | 0.965 | 0.961 | 0.973 | 0.972 |
| Panel D: Agreements | | | | | | |
| (Financial) Sanctions | -0.237*** (0.069) -21.1% | -0.296*** (0.061) -25.6% | -0.229** (0.090) -20.5% | -0.401*** (0.064) -33.0% | -0.754*** (0.119) -53.0% | -0.723*** (0.126) -51.5% |
| Log GDP | 0.422*** (0.089) | 0.720*** (0.034) | 0.566*** (0.104) | 0.519*** (0.099) | 0.744*** (0.213) | 0.832*** (0.201) |
| Economic Integration Agreements | 0.039 (0.058) | -0.020 (0.040) | | | | |
| Services Trade Commitments | | | -0.184*** (0.054) | -0.125* (0.069) | | |
| Bilateral Investment Treaties | | | | | 0.566* (0.311) | 0.585** (0.266) |
| Treaties with Investment Provisions | | | | | 0.017 (0.087) | 0.058 (0.074) |
| Observations | 38,997 | 38,997 | 38,997 | 38,997 | 38,253 | 38,148 |
| Pseudo R² | 0.989 | 0.994 | 0.969 | 0.963 | 0.974 | 0.974 |

Notes: PPML estimation. The dependent variable is specified at the top of each column. The unit of observation is a country-month-year pair. Data cover the period from January 2001 through September 2020 in monthly frequency. Time fixed effects and country-specific fixed effects are included but not reported. Robust standard errors (clustered by country) in parentheses. ***, ** and * denote significant at the 1%, 5% and 10% level, respectively.

Table 4: The Effect of (Financial) Sanctions on Trade in Goods/Services/Financial Flows

| | Goods | | Services | | Financial Flows | |
|--|--------------------------------|--------------------------------|-------------------------------|--------------------------------|--------------------------------|--------------------------------|
| | Imports | Exports | Imports | Exports | Inflows | Outflows |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Panel A: Country-Month-Year Dimension | | | | | | |
| (Financial) Sanctions | -0.281*** (0.085) -24.5% | -0.288*** (0.063) -25.0% | -0.299** (0.142) -25.8% | -0.373*** (0.085) -31.1% | -0.447*** (0.148) -36.0% | -0.512*** (0.141) -40.1% |
| Multilateral Resistance Terms | -0.001* (0.000) | 0.017** (0.008) | 0.001*** (0.000) | 0.018 (0.014) | 0.001** (0.000) | 0.114* (0.058) |
| Observations | 38,664 | 40,944 | 38,664 | 40,944 | 38,016 | 40,260 |
| Pseudo R² | 0.987 | 0.989 | 0.965 | 0.961 | 0.973 | 0.972 |
| Panel B: Product/Declarant-Country-Year Dimension | | | | | | |
| (Financial) Sanctions | -0.138 (0.141) | -0.281*** (0.062) -24.5% | 0.037 (0.170) | -0.011 (0.090) | -0.507*** (0.115) -39.8% | -0.672*** (0.126) -48.9% |
| Multilateral Resistance Terms | -0.000 (0.001) | 0.013* (0.007) | 0.001*** (0.000) | -0.028 (0.022) | 0.000 (0.001) | 0.120* (0.068) |
| Observations | 7,533,356 | 15,285,011 | 3,627,018 | 7,868,572 | 831,887 | 924,464 |
| Pseudo R² | 0.964 | 0.966 | 0.941 | 0.943 | 0.970 | 0.969 |
| Panel C: Product-Country-Month-Year Dimension | | | | | | |
| (Financial) Sanctions | -0.158 (0.146) | -0.240*** (0.067) -21.3% | | | | |
| Multilateral Resistance Terms | -0.000 (0.001) | 0.015* (0.009) | | | | |
| Observations | 89,154,706 | 195,003,413 | | | | |
| Pseudo R² | 0.939 | 0.935 | | | | |

Notes: PPML estimation. The dependent variable is specified at the top of each column. The unit of observation is as defined in each panel. Data cover the period from January 2001 through September 2020 in monthly frequency (panels A and C) or annual frequency (panel B). The following sets of fixed effects are included but not reported: month-year and country-specific fixed effects in panel A; product/declarant-year and product/declarant-country fixed effects in panel B; and product/declarant-month-year and product/declarant-country fixed effects in panel C. Robust standard errors (clustered by country) in parentheses. ***, ** and * denote significant at the 1%, 5% and 10% level, respectively.

Table 5: The Effect of (Financial) Sanctions on Trade in Goods/Services/Financial Flows

| | Goods | | Services | | Financial Flows | |
|--|--------------------------------|--------------------------------|--------------------------------|--------------------------------|--------------------------------|--------------------------------|
| | Imports | Exports | Imports | Exports | Inflows | Outflows |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Total Value (PPML) | -0.306*** (0.077) -26.4% | -0.269*** (0.072) -23.6% | -0.374* (0.192) -31.2% | -0.401*** (0.084) -33.0% | -0.694*** (0.157) -50.0% | -0.645*** (0.163) -47.5% |
| Number of Entries (FLEX) | -0.067 (0.049) | -0.362*** (0.091) -30.4% | -0.255*** (0.089) -22.5% | -0.277*** (0.051) -24.2% | -0.382*** (0.122) -31.8% | -0.381*** (0.123) -31.7% |
| Average Value per Entry (PPML) | -0.097 (0.179) | -0.315*** (0.067) -27.0% | -0.273 (0.271) | -0.056 (0.164) | -0.722 (0.546) | -0.613 (0.450) |
| Number of Declarants/Products (FLEX) | -0.067 (0.049) | -0.362*** (0.091) -30.4% | -0.236*** (0.088) -21.0% | -0.271*** (0.056) -23.7% | -0.363*** (0.102) -30.4% | -0.362*** (0.103) -30.4% |
| Number of Industries/ Categories/Asset Classes (FLEX) | -2.656 (4.157) | -0.004 (0.022) | -0.543*** (0.211) -41.9% | -0.183 (0.139) | -1.093*** (0.424) -66.5% | -1.021** (0.443) -64.0% |

Notes: Results in rows one and three were obtained using PPML. Results in rows two, four, and five were obtained using the Santos Silva et al. (2014) Flex estimator. Each cell contains the coefficient from a separate regression; the regression specification is similar to the corresponding column in Table 3. The dependent variable is listed in the first column; the sample is specified at the top of each column. The unit of observation is a country-month-year pair. Data cover the period from January 2001 through September 2020 in monthly frequency. Robust standard errors (clustered by country) in parentheses. ***, ** and * denote significant at the 1%, 5% and 10% level, respectively.

Table 6: Differential Effect of Export Restrictions

| | Goods | | Services | | Financial Flows | |
|--|--------------------------------|--------------------------------|-------------------|--------------------------------|--------------------------------|-------------------------------|
| | Imports | Exports | Imports | Exports | Inflows | Outflows |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| (Financial) Sanctions | 0.017 (0.087) | -0.039 (0.097) | -0.657 (0.421) | -0.393*** (0.140) -32.5% | -0.822*** (0.311) -56.0% | -0.782** (0.321) -54.3% |
| (Financial) Sanctions combined with export restrictions | -0.385*** (0.113) -32.0% | -0.298*** (0.095) -25.8% | 0.421 (0.435) | -0.016 (0.158) | 0.172 (0.337) | 0.189 (0.347) |
| Observations | 51,192 | 51,192 | 52,140 | 52,140 | 48,822 | 48,822 |
| Pseudo R² | 0.987 | 0.989 | 0.967 | 0.963 | 0.972 | 0.972 |

Notes: PPML estimation. The dependent variable is specified at the top of each column. The unit of observation is a country-month-year pair. Data cover the period from January 2001 through September 2020 in monthly frequency. Time fixed effects and country-specific fixed effects are included but not reported. Robust standard errors (clustered by country) in parentheses. ***, ** and * denote significant at the 1%, 5% and 10% level, respectively.

Table 7: The Effect of (Financial) Sanctions and Possible Anticipation and Removal Effects

| | Goods | | Services | | Financial Flows | |
|---|--------------------------------|--------------------------------|--------------------------------|--------------------------------|--------------------------------|--------------------------------|
| | Imports | Exports | Imports | Exports | Inflows | Outflows |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| (Financial) Sanctions | -0.319*** (0.079) -27.3% | -0.270*** (0.072) -23.7% | -0.403** (0.199) -33.2% | -0.389*** (0.085) -32.2% | -0.688*** (0.157) -49.7% | -0.640*** (0.163) -47.3% |
| Pre-(Financial) Sanctions Period | -0.191 (0.208) | -0.246*** (0.055) -21.8% | -0.215 (0.178) | -0.304** (0.137) -26.2% | -0.736*** (0.281) -52.1% | -0.355 (0.268) |
| (Financial) Sanctions Removed | -0.698*** (0.227) -50.2% | 0.017 (0.226) | -1.334*** (0.474) -73.7% | 0.299 (0.335) | 0.481 (0.447) | 0.358 (0.534) |
| Observations | 51,192 | 51,192 | 52,140 | 52,140 | 48,822 | 48,822 |
| Pseudo R² | 0.987 | 0.989 | 0.967 | 0.963 | 0.972 | 0.972 |

Notes: PPML estimation. The regression specification is similar to the corresponding column in Table 3. The dependent variable is specified at the top of each column. The unit of observation is a country-month-year pair. Data cover the period from January 2001 through September 2020 in monthly frequency. Robust standard errors (clustered by country) recorded in parentheses. ***, ** and * denote significant at the 1%, 5% and 10% level, respectively.

Appendix

Our data on merchandise trade contains information on German imports and exports of some 15,000 different Combined Nomenclature (CN) product codes. This is a much larger number than some 9,500 CN product codes that Eurostat's documentation of the Comext database states are covered in the database. The discrepancy between the two numbers is due to CN codes being periodically revised with some codes discontinued and new codes introduced. Such revision will inflate the number of codes over time. There are typically two ways of dealing with such code revision issues. One is to concord all the changes to CN codes over time and create a synthetic code that captures all the codes that were affected by revisions using an algorithm similar to the one Pierce and Schott (2012) created for U.S. Harmonized System (HS) product classification. The other is to conduct the analysis only on the sample of product codes that were never revised during the sample period and can be thought of having been consistent during the sample period. We now examine our result following the latter approach.

Table A1 compares the results of estimating specification (1) using the full, benchmark sample and the consistent codes sample which drops all CN product codes that changed or were introduced during the sample period. As is readily seen, results for the consistent codes sample are statistically significant and somewhat larger with sanctions reducing imports by 34 percent and exports by 28 percent. In Table A2 we compare the result from estimating specification (2) on both samples with very similar results. We can again conclude that only sanctions with export restrictions reduce imports and exports of goods, but with a larger effect, reducing imports by 39 percent and exports by 30 percent. Note that the number of observations in both samples in Tables A1 and A2 is the same. This is because both samples are created by aggregating the CN-level data to the country level with aggregation obscuring the number of product codes.

While we do not reproduce all tables using the consistent codes sample to conserve space, the general pattern in results is similar to these two tables. In the consistent codes sample estimated coefficients are of the same direction and significance and are somewhat larger. The remaining tables are available on request.

Table A1: The Effect of (Financial) Sanctions on Trade in Goods/Services/Financial Flows

| | Benchmark Sample | | Consistent Codes Sample | |
|------------------------------|--------------------------------|--------------------------------|--------------------------------|--------------------------------|
| | Imports | Exports | Imports | Exports |
| | (1) | (2) | (3) | (4) |
| (Financial) Sanctions | -0.306*** (0.077) -26.4% | -0.269*** (0.072) -23.6% | -0.414*** (0.085) -33.9% | -0.327*** (0.078) -27.9% |
| Observations | 51,192 | 51,192 | 51,192 | 51,192 |
| Pseudo R² | 0.987 | 0.989 | 0.985 | 0.989 |

Notes: PPML estimation. The dependent variable is specified at the top of each column. The unit of observation is a country-month-year pair. Data cover the period from January 2001 through September 2020 in monthly frequency. Time fixed effects and country-specific fixed effects are included but not reported. Robust standard errors (clustered by country) in parentheses. ***, ** and * denote significant at the 1%, 5% and 10% level, respectively.

Table A2: Differential Effect of Export Restrictions

| | Benchmark Sample | | Consistent Codes Sample | |
|---|--------------------------------|--------------------------------|--------------------------------|--------------------------------|
| | Imports | Exports | Imports | Exports |
| | (1) | (2) | (3) | (4) |
| (Financial) Sanctions | 0.017 (0.087) | -0.039 (0.097) | 0.014 (0.143) | -0.053 (0.095) |
| (Financial) Sanctions with export restrictions | -0.385*** (0.113) -32.0% | -0.298*** (0.095) -25.8% | -0.496*** (0.159) -39.1% | -0.356*** (0.096) -30.0% |
| Observations | 51,192 | 51,192 | 51,192 | 51,192 |
| Pseudo R² | 0.987 | 0.989 | 0.985 | 0.989 |

Notes: PPML estimation. The dependent variable is specified at the top of each column. The unit of observation is a country-month-year pair. Data cover the period from January 2001 through September 2020 in monthly frequency. Time fixed effects and country-specific fixed effects are included but not reported. Robust standard errors (clustered by country) in parentheses. ***, ** and * denote significant at the 1%, 5% and 10% level, respectively.